



Williams College ECON 523:

Program Evaluation for International Development

**Lecture 6: Treatment-on-the-Treated**

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photo: Daniela Van Leggelo, Pixella / World Bank

Imperfect Compliance

## How High Is Program Take-Up?

Even “free” programs involve opportunity costs for participants, so take-up is often low

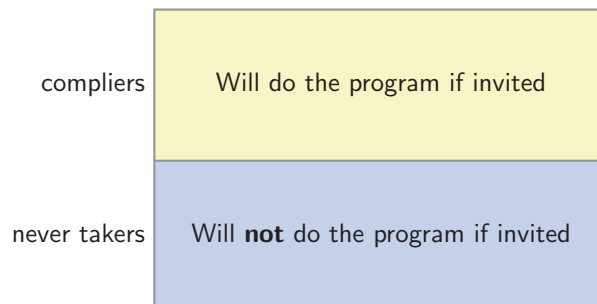
Intervention	Take-Up	Source
Business training	65%	McKenzie & Woodruff (2013)
Deworming medication	75%	Kremer & Miguel (2007)
Microfinance	13% – 31%	JPAL & IPA (2015)

It is often the case that only people who do a program can be impacted by the program\*

- ⇒ We might like to know how much program impacted program **participants**
- ⇒ Not only relevant in randomized trials (who benefits from free primary education?)

\*Often the case, but not always!

## Compliers vs. Never Takers



## Compliers vs. Never Takers

	$T = 0$	$T = 1$
compliers	$Y_{0i}$	$Y_{1i}$
never takers	$Y_{0i}$	$Y_{0i}$

## Imperfect Compliance: A Thought Experiment



Questions:

- What can we say about the average **impact of treatment** on program participants?
- What can we say about the average **outcome** among those who did the program?

## Imperfect Compliance

Suppose outcomes are impacted by program participation ( $P_i$ ), not treatment status ( $T_i$ ):

$$Y_i = Y_{0i} + \delta_i P_i$$

- Program take-up is endogenous conditional on treatment:  $E[Y_{0i}|P_i = 1] \neq E[Y_{0i}|P_i = 0]$
- Only those randomly assigned to treatment ( $T_i = 1$ ) are eligible:  $E[P_i|T_i = 0] = 0$
- Not everyone participates:  $E[P_i|T_i = 1] = \lambda < 1$

Two possible regressions:

- Regress  $Y$  on  $P$  using data from the treatment ( $T_i = 1$ ) group
- Regress  $Y$  on  $T$  using data from the treatment and comparison groups

## How Not to Estimate the Impact of Treatment on the Treated

If we estimate the regression equation  $Y_i = \alpha + \beta P_i + \varepsilon_i$  using data from the treatment group:

$$\begin{aligned}\hat{\beta} &= E[Y_i|P_i = 1] - E[Y_i|P_i = 0] \\ &= E[Y_{1i}|P_i = 1] - E[Y_{0i}|P_i = 0] \\ &= E[Y_{0i} + \delta_i|P_i = 1] - E[Y_{0i}|P_i = 0] \\ &= E[\delta_i|P_i = 1] + E[Y_{0i}|P_i = 1] - E[Y_{0i}|P_i = 0] \\ &= \underbrace{E[\delta_i|\text{compliers}]}_{\text{impact of TOT}} + \underbrace{E[Y_{0i}|\text{compliers}] - E[Y_{0i}|\text{never-takers}]}_{\text{selection bias}}\end{aligned}$$

## The Intent-to-Treat (ITT) Effect

If we estimate the regression equation  $Y_i = \alpha + \beta T_i + \varepsilon_i$ :

$$\hat{\beta} = E[Y_i | T_i = 1] - E[Y_i | T_i = 0]$$

$E[Y_i | T_i = 1]$  is a weighted average of outcomes for complier and never-takers:

$$\begin{aligned} E[Y_i | T_i = 1] &= \lambda E[Y_{1i} | T_i = 1 \text{ and } P_i = 1] + (1 - \lambda) E[Y_{0i} | T_i = 1 \text{ and } P_i = 0] \\ &= \lambda E[\delta_i + Y_{0i} | T_i = 1 \text{ and } P_i = 1] + (1 - \lambda) E[Y_{0i} | T_i = 1 \text{ and } P_i = 0] \\ &= \lambda E[\delta_i | \text{compliers}] + \lambda E[Y_{0i} | \text{compliers}] + (1 - \lambda) E[Y_{0i} | \text{never-takers}] \\ &= \lambda E[\delta_i | \text{compliers}] + E[Y_{0i}] \end{aligned}$$

## The Intent-to-Treat (ITT) Effect

Substituting this into our expression for  $\hat{\beta}$ :

$$\begin{aligned} \hat{\beta} &= E[Y_i | T_i = 1] - E[Y_i | T_i = 0] \\ &= \lambda E[\delta_i | \text{compliers}] + E[Y_{0i}] - E[Y_{0i}] \\ &= \lambda \underbrace{E[\delta_i | \text{compliers}]}_{\text{impact of TOT}} \end{aligned}$$

⇒ Low compliance ( $\lambda < 1$ ) scales down the estimated treatment effect

⇒ ITT effect is average across population ( $T_i = 1$ ), including zero impact on never-takers

## The Impact of Treatment on the Treated

$$\text{ITT} = \lambda \text{TOT} \Leftrightarrow \text{TOT} = \text{ITT}/\lambda$$

The **treatment on the treated (TOT)** estimator:  $\hat{\beta}_{tot} = \frac{E[Y_i|T_i=1] - E[Y_i|T_i=0]}{E[P_i|T_i=1] - E[P_i|T_i=0]}$

- TOT scales up ITT effect to reflect imperfect take-up
- The identifying assumption is that treatment only works through program take-up

## Treatment on the Treated: Implementation (Approach #1)

**Estimating the impact of treatment on the treated via two separate regressions:**

**Intent-to-treat** (aka reduced form): impact of treatment assignment on outcome of interest

$$Y_i = \alpha_{itt} + \beta_{itt} T_i + \varepsilon_i$$

**First stage:** impact of assignment to treatment on program participation:

$$P_i = \alpha_{fs} + \beta_{fs} T_i + \epsilon_i$$

Combine OLS coefficients to estimate impact of treatment on the treated:  $\beta_{tot} = \beta_{itt}/\beta_{fs}$

## Treatment on the Treated: Implementation (Approach #2)

Approach #1 is equivalent to using treatment as an **instrument** for program participation

Assumptions required for instrumental variables estimation:

1. Instrument is exogenous (i.e. not correlated with error term in first stage)
2. Instrument is correlated with treatment (first stage)
3. Only impacts outcomes through program participation (exclusion restriction)

## Treatment on the Treated: Implementation (Approach #2)

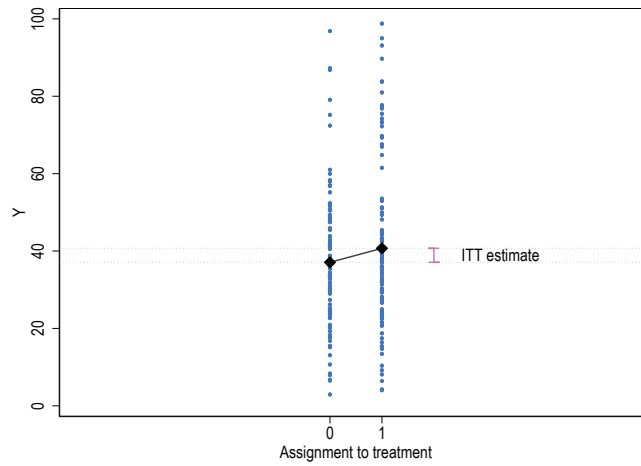
Estimated via two-stage least squares (2SLS):

- **First stage:**  $P_i = \alpha_{fs} + \beta_{fs} T_i + \epsilon_i$
- **Second stage:**  $Y_i = \alpha_{iv} + \beta_{iv} \hat{P}_i + \zeta_i$

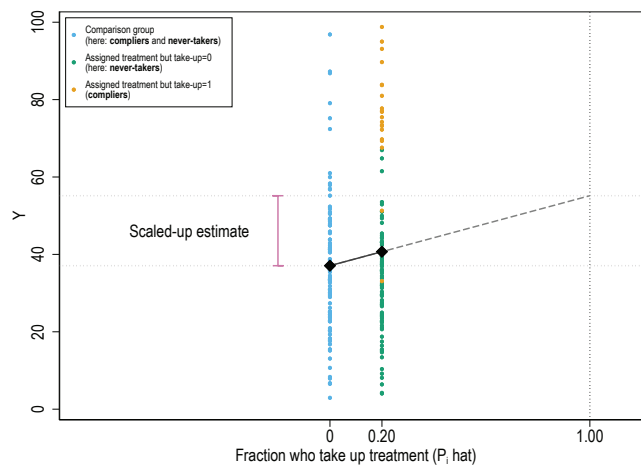
Easy to implement using Stata's **ivregress 2sls** command

- Running two (separate) regressions yields incorrect standard error

## Two-Stage Least Squares (2SLS)



## Two-Stage Least Squares (2SLS)





## Treatment on the Treated: Implementation (Approach #3)

2SLS is also equivalent to a **control function** approach:

- **First stage:**  $P_i = \alpha_{fs} + \beta_{fs} T_i + \epsilon_i$
- **Control function second stage:**  $Y_i = \alpha_{iv} + \beta_{iv} P_i + \gamma \hat{\epsilon}_i + \zeta_i$

First-stage residual captures the endogenous portion of program participation

- Variation in  $P_i$  that remains is the variation explained by  $T_i$
- Second regression equivalent to regressing  $Y_i$  on residuals from a regression of  $P_i$  on  $\hat{\epsilon}_i$

## Treatment on the Treated: Summary of Approaches

1. Divide ITT effect by first stage (impact of  $T$  on  $P$ )
2. Two-stage least squares (regress  $Y$  on predictions from regression of  $P$  on  $T$ )
3. Control function approach (control for residuals from regression of  $P$  on  $T$ )

## Treatment on the Treated: Example

Data from a youth entrepreneurship intervention targeting young women in Nairobi, Kenya

- `treatment` is a dummy for being randomly assigned to the treatment group
- `training` is a dummy for attending at least one day of business training
- `strata` is an ID number for randomization strata (neighborhood×month)
- `income` is a measure of weekly income two years after treatment (from endline survey)

First stage, reduced form regressions take standard form

- **First stage:** regress `training` `treatment` `i.strata`, `r`
- **Reduced form:** regress `income` `treatment` `i.strata`, `r`

## TOT Example: First Stage and Reduced Form Results

	(1)	(2)
	Training	Income
Treatment	0.6105267 (0.0260283) [0.000]	165.9126 (73.81483) [0.025]
Strata fixed effects	Yes	Yes
R-squared	0.470	0.030
Obs.	680	680

Robust standard errors in parentheses; p-values in square brackets.

## TOT Example: Two-Stage Least Squares (2SLS)

Stata syntax for 2SLS:

```
ivregress 2sls income (training = treatment) i.strata, r
```

Generates same coefficients as two-step process, but difference standard errors

```
regress training treatment i.strata, r
predict phat, xb
regress income phat i.strata, r
```

## TOT Example: Two-Stage Least Squares (2SLS)

```
. ivregress 2sls income (training = treatment) i.strata, r
```

income		Robust					
	Coefficient	std. err.	z	P> z	[95% conf. interval]		
training	271.7533	119.5059	2.27	0.023	37.52603	505.9805	
strata							
494002011	243.1708	144.5925	1.68	0.093	-40.22521	526.5668	
494004004	-89.89336	109.9156	-0.82	0.413	-305.324	125.5373	
494004011	39.53772	151.3919	0.26	0.794	-257.185	336.2604	
594004004	52.2759	155.0265	0.34	0.736	-251.5705	356.1222	
594004011	-106.3099	130.9806	-0.81	0.417	-363.0272	150.4073	
594012004	238.6223	146.6926	1.63	0.104	-48.88987	526.1345	
594012011	319.2648	185.929	1.72	0.086	-45.14938	683.6789	
694002004	-167.3286	166.5964	-1.00	0.315	-493.8515	159.1944	
694002011	-187.3286	160.601	-1.17	0.243	-502.1007	127.4436	
694004004	-151.1399	194.2218	-0.78	0.436	-531.8076	229.5278	
694004011	-260.9	196.4015	-1.33	0.184	-645.8398	124.0398	
694012004	209.9024	175.767	1.19	0.232	-134.5947	554.3994	
694012011	233.7189	142.9428	1.64	0.102	-46.44377	513.8815	
_cons	413.216	77.32459	5.34	0.000	261.6626	564.7694	

Instrumental variables 2SLS regression

Number of obs = 680  
Wald chi2(14) = 28.49  
Prob > chi2 = 0.0122  
R-squared = 0.0305  
Root MSE = 950.84

Instrumented: training  
Instruments: 494002011.strata 494004004.strata 494004011.strata  
594004004.strata 594004011.strata 594012004.strata  
594012011.strata 694002004.strata 694002011.strata  
694004004.strata 694004011.strata 694012004.strata  
694012011.strata treatment

## TOT Example: Two-Stage Least Squares (2SLS)

```

. quietly regress training treatment i.strata, r
. predict phat, xb
. regress income phat i.strata, r

```

Linear regression

Number of obs	=	680
F(14, 665)	=	1.97
Prob > F	=	0.0177
R-squared	=	0.0295
Root MSE	=	962

income	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]
phat	271.7533	120.9035	2.25	0.025	34.35466 509.1519
strata					
494002011	243.1708	147.3034	1.65	0.099	-46.06501 532.4066
494004004	-89.89336	112.5547	-0.80	0.425	-310.8988 131.1121

## TOT Example: The Control Function Approach

```

. quietly regress training treatment i.strata, r
. predict resid, resid
. regress income training resid i.strata, r

```

Linear regression

Number of obs	=	680
F(15, 664)	=	1.98
Prob > F	=	0.0144
R-squared	=	0.0322
Root MSE	=	961.36

income	Coefficient	Robust std. err.	t	P> t	[95% conf. interval]
training	271.7533	120.8254	2.25	0.025	34.50743 508.9991
resid	-120.5366	173.7454	-0.69	0.488	-461.6932 220.6199
strata					
494002011	243.1708	146.5754	1.66	0.098	-44.63639 530.978
494004004	-89.89336	111.6222	-0.81	0.421	-309.0684 129.2816

## TOT Example: Interpretation

The entrepreneurship promotion intervention increases income

- TOT effects are larger than ITT effects (is this always true?)
- Assumption: program has no impact on women who do not participate
  - ▶ When might this be a reasonable assumption?
  - ▶ When might this **not** be a reasonable assumption?
- Which is more policy relevant: the ITT estimates or the TOT estimates?
- Could you estimate the TOT impacts of self-employment? Why or why not?

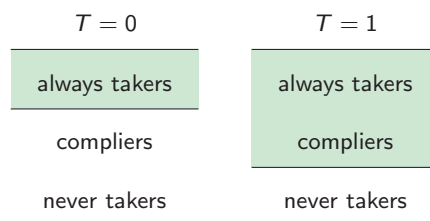
## Two-Sided Non-Compliance

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We sometimes evaluate programs that are available to those in the treatment group

- Examples: medical/health treatment, schooling, vocational/business training, childcare, access to credit, migration, agricultural inputs, management consulting, export contracts
- In such settings, an intervention involves encouraging/facilitating take-up
- Treatment is random and (one hopes) strongly associated with program participation
  - ▶ Compliers participate when assigned to treatment, but not when assigned to control
  - ▶ Some people in the treatment group may choose not to participate
  - ▶ Some people in the control group may still participate in the program

## IV Estimates with Two-Sided Non-Compliance



IV estimates tell us local average treatment effect (LATE) on **compliers**

- Monotonicity assumption: there are no **defiers**
- We can't estimate impacts on **always takers** and **never takers** because being assigned to treatment doesn't change their take-up (i.e. program participation) decision

## Assumptions Required for IV Estimation of LATE

1. Instrument is exogenous (OK in an RCT)
2. Instrument is correlated with treatment (first stage)
3. Only impacts outcomes through take-up (exclusion restriction)
4. Monotonicity (i.e. no defiers)
  - ▶ Treatment either moved people into participation or out of participation, not both
  - ▶ Not required if treatment effects are homogeneous

## Characteristics of the Compliers

The impact of treatment on program participation indicates the proportion compliers

$$E[P_i | T_i = 1] - E[P_i | T_i = 0] = \frac{\text{number of compliers}}{N} = \frac{C}{N}$$

This is also true in sub-populations, e.g. among observations with  $X = 1$  for some  $X$

$$E[P_i | T_i = 1 \text{ and } X_i = 1] - E[P_i | T_i = 0 \text{ and } X_i = 1] = \frac{C_{X=1}}{N_{X=1}}$$

Relative frequency of characteristics  $X = 1$  among compliers, relative to entire population:

$$\frac{E[P_i | T_i=1 \text{ and } X_i=1] - E[P_i | T_i=0 \text{ and } X_i=1]}{E[P_i | T_i=1] - E[P_i | T_i=0]}$$